Analysis of variance and covariance Repeated Measurement Model

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Huda Zeki Naji, Renna Dawood and Rana Rony Faik

Dept.of Mathematics, College of Science, University of Basrah Basrah –Iraq ((Received 14/11/2010, Accepted 16/1/2011))

Abstract

In this article we consider one-way Multivariate repeated measurements analysis of covariance model (MRM ANCOVA), which contains one between-units factor (Group with q levels), one within-units factor (Time with p levels) and one covariate (Z_1) . For this model the covariate is time-independent, that is measured only once. The test statistics of various hypotheses on between-unites factors, within-units factors and the interaction between them are given.

Key Words: (One-Way MRM): One-Way Multivariate Repeated Measures Model, (Λ_r) : Wilks distribution, (U^*) is $P \times P$ orthogonal matrix, the multivariate-Wishart distribution, (ANCOVA) analysis of variance and covariance contain the covariate (Z_1) .

Introduction:

Repeated measurements analysis is widely used in many fields, for example, the health and life science, epidemiology, biomedical, agricultural, industrial, psychological, educational research and so on (see, Huynh and Feldt (1970)[10], and Vonesh and Chinchilli (1997)[8]). Repeated measurements analysis of variance, often referred to as randomized block and split-plot designs (see Bennett and Franklin (1954)[7], Sendeecor and Cochran (1967)[9])."Repeated measurements" is a term used to describe data in which the response variable for each experimental unit is observed on multiple occasions and possibly under different experimental conditions (see,Vonesh and Chinchilli (1997)[8]). This thesis is devoted to study of one-way Multivariate repeated

measurements analysis of covariance model (MRM ANCOVA), which contains one between-units factor (Group with q levels) ,one within-units factor (Time with p levels) and one covariate (Z_1) .For such model, the observations are transformed by an orthogonal matrix to Minimize the probability of presence a correlation between senses of a level of the experiment [3]. The ANCOVA which is based on the first set of transformed observations provides the ANCOVA for the between - units factor effects, while ANCOVA which is based on the k^{th} set of transformed observations, for each k=2,3,...,p provides the ANCOVA for within – units effect [4].

(1.1) Covariate in One-Way Multivariate Repeated Measurements Design :

There is a variety of possibilities for the between units factors in a one-way design. In a randomized one-way MRM experiment, the experimental units are randomized to one between-units factor (Group with q levels) , one within-units factor (Time with p levels) and one covariate $(Z_1)[3]$. For this model the

covariate is time-independent, that is measured only once. For convenience ,we define the following linear model and parameterization for the one-way repeated measurements design with one between units factor incorporation two covariates :-

$$Y_{ijk} = \mu + \tau_j + (\delta_i)_j + \gamma_k + (\tau\gamma)_{jk} + (Z_{1ij} - \overline{Z}_{1..})\beta_1 + e_{ijk} \qquad \dots (1.1)$$

Where

 $(i = 1,...,n_j)$ is an index for experimental unit within group (j)

(j = 1,...,q) is an index for levels of the between-units factor (Group).

(k = 1,..., p) is an index for levels of the within-units factor (Time).

 $Y_{ijk} = (Y_{ijk1},...,Y_{ijkr})'$ is the response measurement of within-units factors (Time) for unit i within treatment factors (Group).

 $\mu = (\mu_1, ..., \mu_r)'$ is the over all mean.

 $\tau_j = (\tau_{j1}, ..., \tau_{jr})'$ is the added effect of the j^{th} level of the treatment factor (Group).

 $\delta_{i(j)} = (\delta_{i(j)1}, ..., \delta_{i(j)r})'$ is the random effect due to the experimental unit i within treatment group j.

 $\gamma_k = (\gamma_{k1}, ..., \gamma_{kr})'$ is the added effect of the k^{th} level of Time.

 $(\tau\gamma)_{jk} = ((\tau\gamma)_{jk1},...,(\tau\gamma)_{jkr})'$ is the added effect of the interaction between the units factor (Group) at the level of (Time).

 $Z_{1ij} = (Z_{1ij1}, ..., Z_{1ijr})'$ is the value of covariate Z_1 at time k for unit i within group j.

 $\overline{Z}_{1..} = (\overline{Z}_{1..1}, ..., \overline{Z}_{1..r})'$ is the mean of covariate Z_1 over all experimental units.

 $\beta_1 = (\beta_{11}, ..., \beta_{1r})'$ is the slope corresponding to covariate Z_1 .

 $e_{ijk} = (e_{ijk1}, ..., e_{ijkr})'$ is the random error at time k for unit i within group j.

For the parameterization to be of full rank, we impose the following set of conditions :

q

$$\sum_{k=1}^{p} \gamma_{k} = 0 \quad , \sum_{j=1}^{q} \tau_{j} = 0$$

$$\sum_{j=1}^{q} (\tau \gamma)_{jk} = 0 \quad \text{for each} \quad k = 1, ..., p \quad , \sum_{k=1}^{p} (\tau \gamma)_{jk} = 0 \quad \text{for each} \quad j = 1, ...,$$

$$\sum_{i,j=1}^{n_{j},q} (Z_{1ij}) = 0$$

We assume that $e_{ijk}^{,s}$ and $\delta_{i(j)}^{,s}$ are independent with $e_{ijk} = (e_{ijk1},..., e_{ijkr})' \sim i.i.d N_r(0, \Sigma_e)$ $\delta_{i(j)} = [\delta_{i(j)1},...,\delta_{i(j)r}]' \sim i.i.d N_r(0, \Sigma_{\delta})$ Where N_r is denoted to the multivariate – Wr normal distribution, and $\Sigma_e, \Sigma_{\delta}$ are $r \times r$ the positive definite matrices [2].

...(1.2)

Where the variance and covariance matrix \sum of the model (1.1) satisfy the assumption of compound symmetry, i.e.

$$i.e \quad \Sigma = I_{p} \otimes \Sigma_{e} + J_{p} \otimes \Sigma_{\delta}$$

$$= \begin{bmatrix} \Sigma_{e} + \Sigma_{\delta} & \Sigma_{\delta} & \cdots & \Sigma_{\delta} \\ \Sigma_{\delta} & \Sigma_{e} + \Sigma_{\delta} & \cdots & \Sigma_{\delta} \\ \vdots & \vdots & \vdots & \vdots \\ \Sigma_{\delta} & \Sigma_{\delta} & \cdots & \Sigma_{e} + \Sigma_{\delta} \end{bmatrix}$$
...(1.3)

Where

operation of two matrices.

 \mathbf{j}_{p} denotes the $P \times 1$ vector of one's.

 J_p denotes the $P \times P$ matrix of one's. and

 I_p denotes the $P \times P$ identity matrix.

 \otimes be the Kroneker product

(1.2) Transforming the one-way Repeated measurements Analysis of Covariance (ANCOVA) model :

In this section we use an orthogonal matrix to transform the observations

 Y_{ijk} for $i = 1,...,n_j$, j = 1,...,q, k = 1,...,p

Let U^* be any $P \times P$ orthogonal matrix and U denotes $p \times (p-1)$ matrix, which is partitioned as follows [4]:

$$U^* = (p^{-\frac{1}{2}} j_p \quad U) \qquad \dots (1.4)$$

$$Y_{ij}^{*} = Y_{ij}U^{*}$$

$$[Y_{ij}^{*} \cdots Y_{ijp}^{*}] = [Y_{ij1} \cdots Y_{ijp}]'U^{*}$$

$$\operatorname{cov}(\vec{Y}_{ij}^{*}) = (U^{*} \otimes I_{r})(I_{p} \otimes \Sigma_{e} + J_{p} \otimes \Sigma_{\delta})(U^{*} \otimes I_{r})$$

$$= I_{p} \otimes \Sigma_{e} + U^{*}J_{p}U^{*} \otimes \Sigma_{\delta} \qquad \dots \qquad 0$$

$$= \begin{bmatrix} \Sigma_{e} + p\Sigma_{\delta} & 0 & \cdots & 0 \\ 0 & \Sigma_{e} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \Sigma_{e} \end{bmatrix}$$
...(1.5)

(1.3) Analysis of Covariance (ANCOVA) for the One-Way Repeated Measurements Model :

In this section, we study the ANCOVA for the effects of between-units factors and within-units factors for the One-way RM model (1.1). Also we give the null hypotheses which are

concerned with these effects and the interaction between them, and the test statistics for them [8]

Now

$$Y_{ij1}^{*} = Y_{ij} p^{-\frac{1}{2}} j_{p} = \left[\frac{1}{\sqrt{p}} \sum_{k=1}^{p} Y_{ijk1} - \frac{1}{\sqrt{p}} \sum_{k=1}^{p} Y_{ijk2} \cdots \frac{1}{\sqrt{p}} \sum_{k=1}^{p} Y_{ijkr} \right]^{r}$$

From (1.1), we obtain [3]:

$$Y_{ij1}^{*} = \mu^{*} + \tau_{j}^{*} + \delta_{i(j)}^{*} - \overline{Z}_{1..}^{*}\beta_{1}^{*} + e_{ij1}^{*}$$

Then the set of vectors
 $(Y_{111}^{*}, \dots, Y_{n_{1}11}^{*})', (Y_{121}^{*}, \dots, Y_{n_{2}21}^{*})', \dots, (Y_{1q1}^{*}, \dots, Y_{n_{q}q1}^{*})'$
Have mean vectors [11] :
 $X_{1} = \sqrt{P}\mu + \sqrt{P}\tau_{1} - \sqrt{P}\overline{Z}_{1..}\beta_{1}$
 $X_{2} = \sqrt{P}\mu + \sqrt{P}\tau_{2} - \sqrt{P}\overline{Z}_{1..}\beta$
:
 $Y_{2} = \sqrt{P}\mu + \sqrt{P}\tau_{2} - \sqrt{P}\overline{Z}_{1..}\beta$

$$X_{q} = \sqrt{P}\mu + \sqrt{P}\tau_{q} - \sqrt{P}\overline{Z}_{1..}\beta_{1..}\beta_{1..}\beta_{1..}\beta_{1....}\beta_{1...}\beta_{1...}\beta_{1...}\beta_{1...}\beta_{1$$

Respectively, and each of them has covariance matrix $p \sum_{\delta} + \sum_{e}$. So, the null hypothesis of the same treatment effects is [1]:

$$H_{01} = \tau_1 - \overline{Z}_{1..} = \dots = \tau_q - \overline{Z}_{1..} = 0$$

The ANCOVA based on the set of transformed to the following form for the sum square terms observations above the $Y_{ij1,s}^*$ provides the SS_G , SS_{Z_1} , $SS_{u(GroupZ_1)}[3]$: ANCOVA for between-units effects. This leads

$$SS_{G} = \sum_{j=1}^{q} n_{j} (\overline{Y}_{j1}^{*} - \overline{Y}_{1}^{*}) (\overline{Y}_{j1}^{*} - \overline{Y}_{1}^{*})', \quad \overline{Y}_{j1}^{*} = \frac{\sum_{i=1}^{n_{j}} Y_{ij1}^{*}}{n_{j}}, \quad \overline{Y}_{1}^{*} = \frac{\sum_{j=1}^{q} \sum_{i=1}^{n_{j}} Y_{ij1}^{*}}{n}$$

$$SS_{Z_{1}} = \sum_{j=1}^{q} \sum_{i=1}^{n_{j}} \frac{1}{\beta_{1}} \Big[(\overline{Y}_{ij1}^{*} - \overline{Y}_{j1}^{*}) (\overline{Y}_{ij1}^{*} - \overline{Y}_{j1}^{*})' \Big]$$

$$SS_{u(GroupZ_{1})} = \sum_{j=1}^{q} \sum_{i=1}^{n_{j}} \left((Y_{ij1}^{*} - \overline{Y}_{ij.}^{*} + \overline{Y}_{1}) (Y_{ij1}^{*} - \overline{Y}_{ij.}^{*} + \overline{Y}_{1})' \right)$$

Thus

$$\begin{split} SS_G &\sim W_r(q-1, p\sum_{\delta} + \sum_e) \\ SS_{Z_1} &\sim W_r(1, p\sum_{\delta} + \sum_e) \\ SS_{u(GroupZ_1Z_2)} &\sim W_r(n-q-2, p\sum_{\delta} + \sum_e) \end{split}$$

Where W_r denotes the multivariate-Wishart distribution. The multivariate Wilks test [12] :

$$\begin{split} T_{w_1} &= \frac{\left|SS_{u(GroupZ_1)}\right|}{\left|SS_{u(GroupZ_1)} + SS_G\right|} \quad \text{, When } H_{01} \text{ is true} \\ T_{w_1} &\sim \Lambda_r \left(n - q - 2, q - 1\right) \\ T_{w_2} &= \frac{\left|SS_{u(GroupZ_1)}\right|}{\left|SS_{u(GroupZ_1)} + SS_{Z_1}\right|} \quad \text{, When } H_{01} \text{ is true} \\ T_{w_2} &\sim \Lambda_r \left(n - q - 2, 1\right) \end{split}$$

The ANCOVA based on the set of transformed observations the Y_{ijk}^* for each k = 2, ..., p has the model which is partitioned as follows[3]:

$$Y_{ijk}^{*} = \sum_{k'=1}^{p} u_{kk'} Y_{ijk'}, \quad k = 2, \cdots, p$$

$$= \sum_{k'=1}^{p} u_{kk'} \left(\mu + \tau_{j} + \gamma_{k'} + (\tau\gamma)_{jk'} + (Z_{1ij} - \overline{Z}_{1..})\beta_{1} + (Z_{2ij} - \overline{Z}_{2..})\beta_{2} + e_{ijk'} \right)$$

$$Y_{ijk}^{*} = \gamma_{k}^{*} + (\tau\gamma)_{jk}^{*} - \overline{Z}_{1..}^{*}\beta_{1}^{*} + e_{ijk}^{*}$$

Then from above the analysis we test the following hypotheses[4]:

$$H_{02} : \gamma_{2}^{*} - \overline{Z}_{2..}^{*} = \dots = \gamma_{p}^{*} - \overline{Z}_{1..}^{*} = 0$$
$$H_{03} : (\tau\gamma)_{j2}^{*} - \overline{Z}_{1..}^{*} = \dots = (\tau\gamma)_{jp}^{*} - \overline{Z}_{1..}^{*} = 0$$

The ANCOVA based on the set of transformed observations above , the $Y_{ij2}^*, \ldots, Y_{ijp}^*$ provides the ANCOVA for within-units effects.

This leads to the following forms for the sum square terms[3]:

$$SS_{Time} = n \sum_{k=2}^{p} (\overline{Y}_{k}^{*} (\overline{Y}_{k}^{*})')$$

$$SS_{Time \times Group} = \sum_{k=2}^{p} \sum_{j=1}^{q} n_{j} ((\overline{Y}_{jk}^{*} - \overline{Y}_{k}^{*})(\overline{Y}_{jk}^{*} - \overline{Y}_{k}^{*})'), \quad \overline{Y}_{jk}^{*} = \frac{\sum_{i=1}^{n_{j}} Y_{ijk}^{*}}{n_{j}} , k = 2, ..., p$$

$$SS_{Time \times Umit(GroupZ_{1})} = \sum_{k=2}^{p} \sum_{j=1}^{q} \sum_{i=1}^{n_{j}} ((Y_{ijk}^{*} - \overline{Y}_{ij.} + \overline{Y}_{k})(Y_{ijk}^{*} - \overline{Y}_{ij.} + \overline{Y}_{k})')$$

Then from the above sum square terms , we have [13] : $SS_{Time} \sim W_r((p-1), p\sum_{\delta} + \sum_e)$ $SS_{Time\times Group} \sim W_r((p-1)(q-1), p\sum_{\delta} + \sum_e)$ $SS_{Time\times Unit(GroupZ_1)} \sim W_r((p-1)(n-q)p\sum_{\delta} + \sum_e)$ The multivariate Wilks test [12] :

$$T_{w_3} = \frac{\left|SS_{Time \times Unit(GroupZ_1)}\right|}{\left|SS_{Time \times Unit(GroupZ_1)} + SS_{Time}\right|}, \text{ when } H_{02} \text{ is true}$$
$$T_{w_3} \sim \Lambda_r \left((p-1)(n-q), (p-1)\right)$$
$$T_{w_4} = \frac{\left|SS_{Time \times Unit(GroupZ_1)}\right|}{\left|SS_{Time \times Unit(GroupZ_1)} + SS_{Time \times Group}\right|}, \text{ when } H_{02} \text{ is true}$$
$$T_{w_4} \sim \Lambda_r \left((p-1)(n-q), (p-1)(q-1)\right)$$

The one-way MRM ANCOVA with one between-unit factor (Group) and two covariates (Z_1) that are time-independent.

ume-independent.				
	Source	D.F	SS	Wilks Criterion
Between	Group	q - 1	SS _G	$T_{W_1} = \frac{\left SS_{Unit(GroupZ_1)}\right }{\left SS_{Unit(GroupZ_1)} + SS_G\right }$
	Z_1	1	SS _{Z1}	$T_{W_2} = \frac{\left SS_{Unit(GroupZ_1)}\right }{\left SS_{Unit(GroupZ_1)} + SS_{Z_1}\right }$
	Unit(GroupZ ₁)	n-q-2	$SS_{Unit(GroupZ_1)}$	
Within	Time	<i>p</i> – 1	SS_{Time}	$T_{W_{3}} = \frac{\left SS_{Time \times Unit(GroupZ_{1})}\right }{\left SS_{Time \times Unit(GroupZ_{1})} + SS_{Time}\right }$
	Time × Group	$(p-1) \times (q-1)$	$SS_{Time imes Group}$	$T_{W_4} = \frac{\left SS_{Time \times Unit(GroupZ_1)}\right }{\left SS_{Time \times Unit(GroupZ_1)} + SS_{Time \times Group}\right }$
	Time*Unit(GroupZ ₁)	$(p-1) \times (n-q)$	$SS_{Time \times Unit(Group Z_1)}$	

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المستخلص

One- تتاول هذا البحث دراسة نموذج من نماذج القياسات المتكررة متعددة المتغيرات للبيانات الكاملة ذات الاتجاه الواحد (-One way MRM) لتحليل التباين المشترك، والنموذج الذي تم دراسته هنا ذا اتجاه واحد ، وهذا يدل على وجود عامل واحد بين الوحدات يدعى (المجموعة) وعامل واحد داخل الوحدات يدعى الزمن هذا فضلاً عن عامل مرافق (مستقل). كون هذا النموذج يحتوي على عامل مرافق (مستقل) واحد ، لذا يطلق على النموذج العام اسم (ANCOVA) ،ويشترط فى هذا

النموذج أن يكون العامل المرافق) Z1 (ثابت و مستقل زمنيا (أي يقاس مرة واحدة في كل مستوى من مستويات التجربة).